

# New Insights into History Matching via Sequential Monte Carlo

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# Parameter Estimation

Denote  $\theta$  as the parameter of a deterministic model  $M(\theta)$  or stochastic model  $M(\theta, u)$  where  $u$  are the random numbers needed in the simulation.

Have some observed data  $y$  believed to be generated by  $M$ .

Wish to obtain a collection of plausible  $\theta$  based on  $y$ . Eg a Bayesian posterior,  $\pi(\theta|y) \propto p(y|\theta)\pi(\theta)$ .

However, the model  $M$  is too expensive to use say Markov chain Monte Carlo or sequential Monte Carlo.

# History Matching

Aim: Determine 'plausible' region of the parameter space relatively quickly.

History matching (eg Craig et al 1997) **determines a non-implausible region of the parameter space relatively quickly** by using an *emulator* of the model outputs or 'distance' between model outputs and  $y$ .

The **emulator not only provides predictions at untrained points but also quantifies the uncertainty in the predictions.**

# History Matching Notation/Details

Denote  $\pi(\theta)$  as a 'prior' distribution.

For simplicity we will assume there is only  $n = 1$  model output, eg:

- The model really does have only 1 output
- The output is a distance between outputs and data.
- The output is a (approximate) likelihood function.

Parameter  $\theta$  is deemed as **non-implausible** if  $\mathcal{I}(\theta) < c$  for cut-off  $c$  where  $\mathcal{I}(\theta)$  includes emulation uncertainty, eg Andrianakis et al 2015

$$\mathcal{I}(\theta) = \frac{|y_{\theta} - y_{\text{obs}}|}{\sqrt{s_{m,\theta}^2 + s_{e,\theta}^2 + s_d^2}}. \quad (1)$$

Uncertainty:  $s_{m,\theta}^2$  simulator,  $s_{e,\theta}$  emulator,  $s_d^2$  model.

# History Matching Procedure

Steps involved in the history matching algorithm:

- 1 Generate  $N_w$  training samples  $\{\theta_j\}_{j=1}^{N_w} \sim \pi(\theta)$  using a space filling design and simulate the model at each  $\theta_j$  to generate the collection of outputs  $\{y_{\theta_j}\}_{j=1}^{N_w}$ .
- 2 Fit an emulator  $E_w$  to the training data  $\{\theta_j, y_{\theta_j}\}_{j=1}^{N_w}$ .
- 3 Use the emulator  $E_w$  to define an implausibility function  $\mathcal{I}_w(\theta)$ . If  $\mathcal{I}_w(\theta) > c_w$  for some chosen  $c_w$  then  $\theta$  is deemed as implausible by emulator  $E_w$ .
- 4 Use all emulators  $\{E_r\}_{r=1}^w$  to define the non-implausible region  $\Theta_w = \{\theta \in \Theta \mid \bigcap_{r=1}^w \mathcal{I}_w(\theta) < c_w\}$ .
- 5 Increase wave counter  $w = w + 1$ .  
Generate  $N_w$  training samples  $\{\theta_j\}_{j=1}^{N_w}$  from  $\Theta_w$  and simulate the model at each  $\theta_j$  to generate  $\{y_{\theta_j}\}_{j=1}^{N_w}$ .
- 6 If the stopping rule is satisfied then finish otherwise return to Line 2.

# Issues with History Matching

History matching does have at least two issues:

- 1 The cut-off values  $c_w$  may not be easy to select in practice and there is no existing automated method for doing so.
- 2 Sampling uniformly from  $\Theta_w$  as  $w$  increases becomes increasingly difficult. History matching papers tend to “ignore” this.

Solution: We use **sequential Monte Carlo** (SMC, eg Chopin 2002) to help address the above 2 issues.

# Sequential Monte Carlo

SMC samples from a sequence of distributions (connecting easy and target distributions) by iteratively applying re-weighting, re-sampling and move steps.

Advantages of SMC approach over MCMC:

- Naturally adaptive
- Easily parallelisable
- More capable of dealing with multimodal or complex posterior distributions

# SMC History Matching

For history matching, we define the sequence of distributions as

$$p_w(\theta) \propto \pi(\theta) \prod_{k=1}^w \mathbb{I}(\mathcal{I}_k(\theta) \leq c_k).$$

Assume that we have a collection of ‘particles’,  $\{W_w^i, \theta_w^i\}_{i=1}^M$  from  $p_w(\theta)$ . To push the particles to target  $w + 1$  we apply **re-weighting step**:

$$W_{w+1}^i \propto W_w^i \mathbb{I}(\mathcal{I}_{w+1}(\theta_w^i) \leq c_{w+1}).$$

Here  $W_w$  are all equal, so the weights  $W_{w+1}$  will either be constant or equal to zero. Thus we can select  $c_{w+1}$  so that a certain proportion,  $\alpha$ , have non-zero weight. Same as ensuring that the effective sample size (ESS) =  $1 / \sum_{i=1}^M (W_{w+1}^i)^2$ , is  $\alpha M$ .



# SMC History Matching

After re-weighting the ESS drops to roughly  $\alpha M$ .

**Resampling**  $M$  times from the surviving particles allows the ESS to return to  $M$ . However, the drawback is that some particles will be duplicated.

**Diversify particles with an MCMC kernel**  $R$  times to each of the resampled particles. Can adaptively choose  $R$ . **Can use population of particles to inform efficient MCMC proposal**. This move step only uses the emulator (no expensive model simulations).

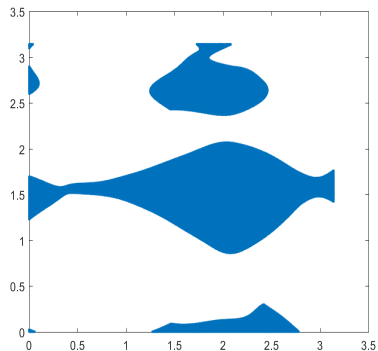
# SMC History Matching – Move Step

History matching produces irregular and multi-modal distribution, need efficient MCMC kernel.

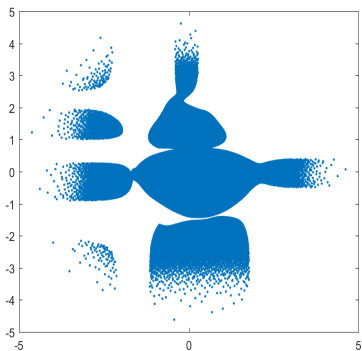
We use a kernel density estimate (kde) of each component. Using the cdf of the kde we transform each margin to roughly  $\mathcal{N}(0, 1)$ . Then we perform a random walk on the transformed space.

Modelling of the marginals can help bring modes closer together and produces a more regular space to perform a random walk.

# Modelling of Marginals



(a) before



(b) after

# Emulation

For emulation we use a Gaussian Process (GP). A 'prior' over functions when evaluated at a finite number of points has a multivariate normal distribution.

Characterised by a mean function  $m_{\beta}(\theta)$  and covariance function  $C_{\gamma}(\theta, \theta') = \text{cov}\{f(\theta), f(\theta')\}$ , with hyperparameters  $\beta$  and  $\gamma$ .

Here we take  $m_{\beta}(\theta) = 0$  and

$$C_{\gamma}(\theta, \theta') = \delta_c \exp \left\{ -\frac{1}{2} \sum_{k=1}^p \frac{(\theta_k - \theta'_k)^2}{r_k^2} \right\},$$

with hyperparameters  $\gamma = (\delta_c, r_1, \dots, r_p)^{\top}$ .

## Emulation (Cont...)

Also include ‘nugget’ term:

$$\text{cov}\{\hat{f}(\boldsymbol{\theta}), \hat{f}(\boldsymbol{\theta}')\} = \mathbf{C}_{\hat{\gamma}}(\boldsymbol{\theta}, \boldsymbol{\theta}') + \delta \mathbf{1}(\boldsymbol{\theta} = \boldsymbol{\theta}')$$

Based on “training” data,  $\Theta$ , hyperparameters estimated via (marginal) maximum likelihood estimation. Obtain “posterior” over functions.

Can use to predict at untrained location and quantify prediction uncertainty:

$$m^*(\boldsymbol{\theta}) = \mathbf{C}_{\hat{\gamma}}(\boldsymbol{\theta}, \Theta)^\top \{ \mathbf{C}_{\hat{\gamma}}(\Theta, \Theta) + \hat{\delta} \mathbf{I} \}^{-1} \mathbf{f}(\Theta),$$
$$s^*(\boldsymbol{\theta})^2 = \mathbf{C}_{\hat{\gamma}}(\boldsymbol{\theta}, \boldsymbol{\theta}) - \mathbf{C}_{\hat{\gamma}}(\boldsymbol{\theta}, \Theta)^\top \{ \mathbf{C}_{\hat{\gamma}}(\Theta, \Theta) + \hat{\delta} \mathbf{I} \}^{-1} \mathbf{C}_{\hat{\gamma}}(\Theta, \boldsymbol{\theta}).$$

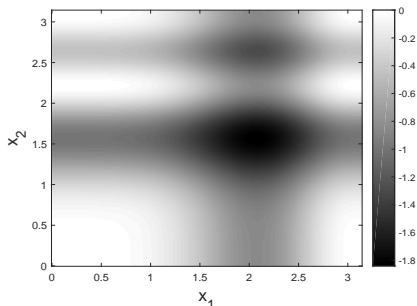
# Simple Example

Consider the function

$$y = -\sin(x_1) \sin(x_1^2/\pi)^2 - \sin(x_2) \sin(2x_2^2/\pi)^2,$$

where  $\theta = (x_1, x_2) \in (0, \pi) \times (0, \pi)$ .

We wish to find the regions where  $y$  is small.



# History Matching Parameters

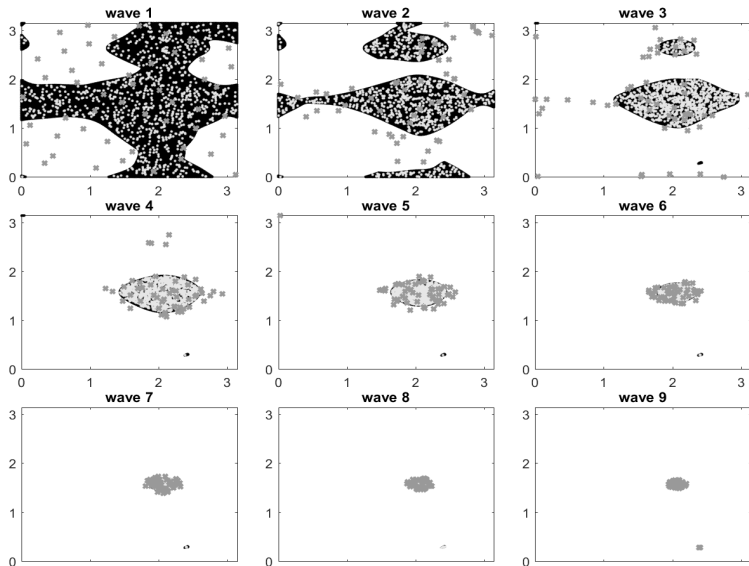
We use  $N = 50$  training samples at each wave.

Implausability measure  $\mathcal{I}(\theta) = y_p(\theta) - r \times s_p(\theta)$  where  $y_p(\theta)$  and  $s_p(\theta)$  is the prediction and the standard deviation from the currently fitted GP, respectively. Set  $r = 3$ .

We solve the history matching problem ‘perfectly’ by taking  $2^{20}$  draws from parameter space and use rejection sampling. The cut-offs  $c_w$  are chosen so that half the remaining particles survive at each wave.

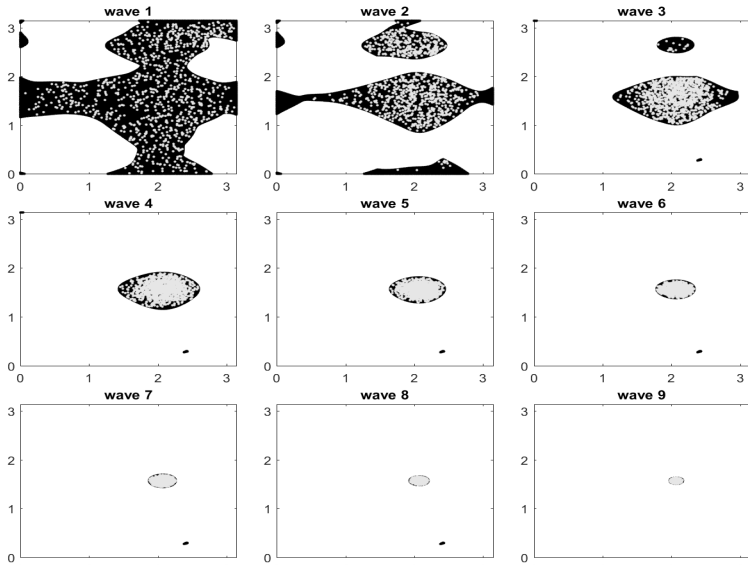
We investigate how well SMC can uniformly sample from non-implausible space implied by ‘perfect’ history matching.

# Results from SMC History Matching





# Results from Typical History Matching Approach



## More Challenging Example

Rainfall Run-off Model in Hydrology. 7 parameters.

Underlying model is deterministic. Output is distance between observed and “simulated” time series

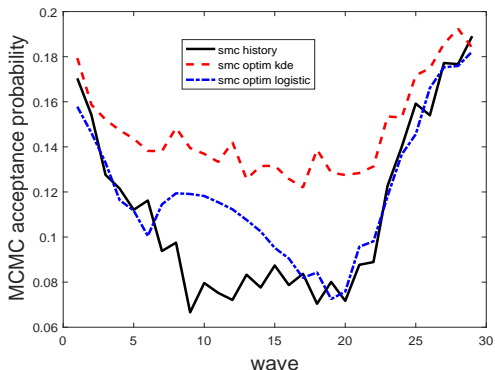
$$\rho_p(\theta) = \sum_{t=1}^T \frac{(y_{\text{obs}}^t - y_{\theta}^t)^2}{y_{\text{obs}}^t},$$

We compare SMC History Matching with standard SMC optimisation (no emulator):

$$p_w(\theta) \propto \pi(\theta) \mathbb{I}(\rho_p \leq d_w), \text{ for } d_1 > d_2 > \dots > d_w > \dots,$$

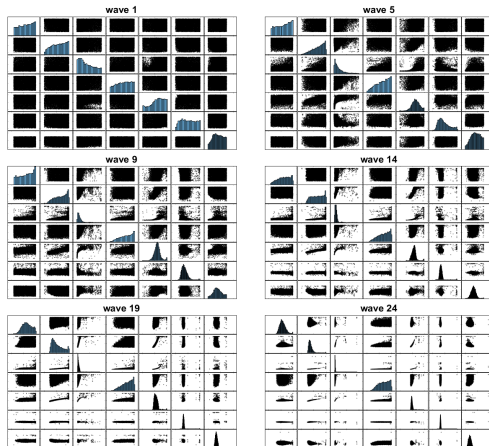
with  $d_w$  determined adaptively based on quantiles of  $\rho_p$ .

# Results



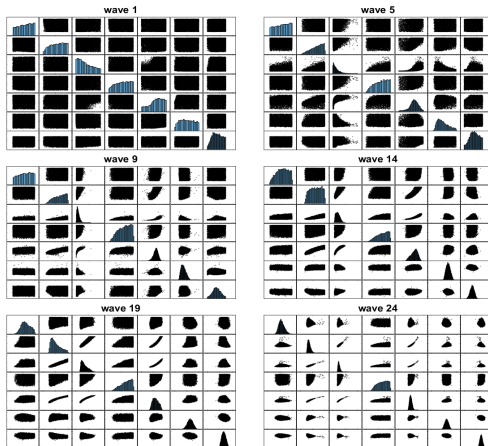
Acceptance probability of the MCMC step in SMC history matching method (solid) and SMC optimisation method (dash).

# Results



Bivariate scatterplots of the parameters with SMC history matching.

# Results



Bivariate scatterplots of the parameters for SMC optimisation

# Results

SMC optimisation has less challenging sampling problem.

But... SMC History Matching uses 2 orders of magnitude less evaluations of expensive simulator.

# Take Home Message

History Matching is useful for finding non-implausible regions for complex simulators quickly. But...

It replaces the hard parameter calibration problem with a hard sampling problem.

But... SMC can help here, and create a more automatic history matching algorithm.

# References

Andrianakis et al (2015). Bayesian history matching of complex infectious disease models using emulation: a tutorial and a case study on HIV in Uganda. PLOS Computational Biology, 11(1):e1003968.

Chopin (2002). A sequential particle filter method for static models. Biometrika, 89(3):539-552.

Craig et al (1997). Pressure matching for hydrocarbon reservoirs: a case study in the use of Bayes linear strategies for large computer experiments. In Case Studies in Bayesian Statistics, pages 37-93. Springer.

Drovandi et al (2017). [New insights into history matching via sequential Monte Carlo.](https://eprints.qut.edu.au/112175/)  
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